# IIT JAM 2018 Mathematical Statistics (MS) Question Paper

Time Allowed: 3 Hours | Maximum Marks: 100 | Total questions: 60

# **General Instructions**

#### **General Instructions:**

- i) All questions are compulsory. Marks allotted to each question are indicated in the margin.
- ii) Answers must be precise and to the point.
- iii) In numerical questions, all steps of calculation should be shown clearly.
- iv) Use of non-programmable scientific calculators is permitted.
- v) Wherever necessary, write balanced chemical equations with proper symbols and units.
- vi) Rough work should be done only in the space provided in the question paper.

1. Let  $\{a\_n\}$  be a sequence of real numbers such that  $a_1=2$ , and for  $n\geq 1$ ,  $a_{n+1}=\frac{2a_n+1}{a_n+1}$ .

- (A)  $1.5 \le a_n \le 2$ , for all natural numbers  $n \ge 1$
- (B) There exists a natural number  $n \ge 1$  such that  $a_n > 2$
- (C) There exists a natural number  $n \ge 1$  such that  $a_n < 1.5$
- (D) There exists a natural number  $n \ge 1$  such that  $a_n = \frac{1+\sqrt{5}}{2}$

2. The value of

$$\lim_{n\to\infty} \left(1+\frac{2}{n}\right)^{n^2} e^{-2n} \text{ is}$$

- (A)  $e^{-2}$
- (B)  $e^{-1}$
- (C) e
- (D)  $e^2$

3. Let  $\{a\_n\}$  and  $\{b\_n\}$  be two convergent sequences of real numbers. For  $n \ge 1$ , define  $u_n = \max\{a_n, b_n\}$  and  $v_n = \min\{a_n, b_n\}$ . Then

- (A) Neither {a\_n} nor {b\_n} converges
- (B)  $\{u_n\}$  converges but  $\{v_n\}$  does not converge
- (C)  $\{u_n\}$  does not converge but  $\{v_n\}$  converges
- (D) Both  $\{u_n\}$  and  $\{v_n\}$  converge

4. Let

$$M = \begin{pmatrix} 1 & 4 \\ 3 & 2 \end{pmatrix}.$$

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If I is the  $2 \times 2$  identity matrix and 0 is the  $2 \times 2$  zero matrix, then

- (A)  $20M^2 13M + 7I = 0$
- (B)  $20M^2 13M 7I = 0$

- (C)  $20M^2 + 13M + 7I = 0$
- (D)  $20M^2 + 13M 7I = 0$

# 5. Let X be a random variable with the probability density function

$$f(x) = \begin{cases} \frac{x^p}{\Gamma(p)} e^{-\alpha x} x^{p-1}, & x \ge 0, \ \alpha > 0, \ p > 0, \\ 0, & \text{otherwise.} \end{cases}$$

If E(X) = 20 and Var(X) = 10, then  $(\alpha, p)$  is

- (A)(2,20)
- (B)(2,40)
- (C)(4,20)
- (D)(4,40)

### **6.** Let X be a random variable with the distribution function

$$F(x) = \begin{cases} 0, & x < 0, \\ \frac{1}{4} + \frac{4x - x^2}{8}, & 0 \le x < 2, \\ 1, & x \ge 2. \end{cases}$$

**Then** 

$$P(X=0) + P(X=1.5) + P(X=2) + P(X \ge 1)$$

equals

- (A)  $\frac{3}{8}$
- (B)  $\frac{5}{8}$
- (C)  $\frac{7}{8}$
- (D) 1

# 7. Let $X_1, X_2$ and $X_3$ be i.i.d. U(0,1) random variables. Then

$$E\left(\frac{X_1 + X_2}{X_1 + X_2 + X_3}\right)$$

## equals

- (A)  $\frac{1}{3}$
- (B)  $\frac{1}{2}$
- (C)  $\frac{2}{3}$
- (D)  $\frac{3}{4}$

8. Let  $x_1 = 0, x_2 = 1, x_3 = 2, x_4 = 3$  and  $x_5 = 0$  be the observed values of a random sample of size 5 from a discrete distribution with the probability mass function

$$f(x;\theta) = P(X = x) = \begin{cases} \frac{\theta}{3}, & x = 0, \\ \frac{2\theta}{3}, & x = 1, \\ \frac{1-\theta}{2}, & x = 2, 3, \end{cases}$$

where  $\theta \in [0,1]$  is the unknown parameter. Then the maximum likelihood estimate of  $\theta$  is

- (A)  $\frac{2}{5}$
- (B)  $\frac{3}{5}$
- (C)  $\frac{5}{7}$
- (D)  $\frac{5}{9}$

9. Consider four coins labelled as 1, 2, 3 and 4. Suppose that the probability of obtaining a 'head' in a single toss of the *i*th coin is  $\frac{1}{4}$ , i=1,2,3,4. A coin is chosen uniformly at random and flipped. Given that the flip resulted in a 'head', the conditional probability that the coin was labelled either 1 or 2 equals

- (A)  $\frac{1}{10}$
- (B)  $\frac{2}{10}$
- (C)  $\frac{3}{10}$
- (D)  $\frac{4}{10}$

## 10. Consider the linear regression model

 $y_i = \beta_0 + \beta_1 x_i + \epsilon_i$ , i = 1, 2, ..., n, where  $\epsilon_i$  are i.i.d. standard normal random variables. Given that

$$\frac{1}{n}\sum_{i=1}^{n} x_i = 3.2, \quad \frac{1}{n}\sum_{i=1}^{n} y_i = 4.2, \quad \frac{1}{n}\sum_{j=1}^{n} \left(x_j - \frac{1}{n}\sum_{i=1}^{n} x_i\right)^2 = 1.5,$$

$$\frac{1}{n}\sum_{i=1}^{n} \left(x_j - \frac{1}{n}\sum_{i=1}^{n} x_i\right) \left(y_j - \frac{1}{n}\sum_{i=1}^{n} y_i\right) = 1.7,$$

the maximum likelihood estimates of  $\beta_0$  and  $\beta_1$ , respectively, are

- (A)  $\frac{17}{15}$  and  $\frac{32}{75}$
- (B)  $\frac{32}{75}$  and  $\frac{17}{15}$
- (C)  $\frac{43}{75}$  and  $\frac{17}{15}$
- (D)  $\frac{43}{75}$  and  $\frac{5}{9}$

# **11.** Let $f:[-1,1] \to \mathbb{R}$ be defined by

 $f(x) = \frac{x^2 + [\sin(\pi x)]}{1 + |x|}$ , where [y] denotes the greatest integer less than or equal to y.

**Then** 

- (A) f is continuous at -1, 0, 1
- (B) f is discontinuous at  $-1, 0, \frac{1}{2}$
- (C) f is discontinuous at  $-1, \frac{1}{2}, 0, \frac{1}{2}$
- (D) f is continuous everywhere except at 0

# **12.** Let $f, g : \mathbb{R} \to \mathbb{R}$ be defined by

$$f(x) = x^2 - \frac{\cos(x)}{2}, \quad g(x) = \frac{x\sin(x)}{2}.$$

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Then

(A) f(x) = g(x) for more than two values of x

- (B)  $f(x) \neq g(x)$ , for all  $x \in \mathbb{R}$
- (C) f(x) = g(x) for exactly one value of x
- (D) f(x) = g(x) for exactly two values of x

13. Consider the domain  $D=\{(x,y)\in\mathbb{R}^2:x\leq y\}$  and the function  $h:D\to\mathbb{R}$  defined by

$$h((x,y)) = (x-2)^4 + (y-1)^4, \quad (x,y) \in D.$$

Then the minimum value of h on D equals

- (A)  $\frac{1}{2}$
- (B)  $\frac{1}{4}$
- (C)  $\frac{1}{8}$
- (D)  $\frac{1}{16}$

14. Let  $M=[X\ Y\ Z]$  be an orthogonal matrix with  $X,Y,Z\in\mathbb{R}^3$  as its column vectors. Then

$$Q = XX^T + YY^T$$
 and  $QZ = Z$ 

implies

- (A) M is a skew-symmetric matrix
- (B) M is the  $3 \times 3$  identity matrix
- (C)  $Q^2 = Q$
- (D) Q satisfies QZ = Z

**15.** Let  $f:[0,3] \to \mathbb{R}$  be defined by

$$f(x) = \begin{cases} 0, & 0 \le x < 1, \\ e^{x^2} - e, & 1 \le x < 2, \\ e^{x^2} + 1, & 2 \le x \le 3. \end{cases}$$

Now, define  $F:[0,3]\to\mathbb{R}$  by

$$F(0) = 0$$
 and  $F(x) = \int_0^x f(t) dt$ , for  $0 < x \le 3$ .

**Then** 

(A) F is differentiable at x = 1 and F'(1) = 0

(B) F is differentiable at x = 2 and F'(2) = 0

(C) F is not differentiable at x = 1

(D) F is differentiable at x = 2 and F''(2) = 1

# 16. If x, y, z are real numbers such that

$$4x + 2y + z = 31$$
 and  $2x + 4y - z = 19$ ,

then the value of 9x + 7y + z is

(A) cannot be computed from the given information

(B) equals  $\frac{281}{3}$ 

(C) equals  $\frac{182}{3}$ 

(D) equals  $\frac{218}{3}$ 

#### 17. Let

$$M = \begin{pmatrix} 1 & 1 & -1 \\ -1 & 1 & 1 \\ -1 & -1 & 1 \end{pmatrix}.$$

If

$$V = \{(x, y, 0) \in \mathbb{R}^3 : M \begin{pmatrix} x \\ y \\ 0 \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \\ 0 \end{pmatrix} \},$$

$$W = \{(x, y, z) \in \mathbb{R}^3 : M \begin{pmatrix} x \\ y \\ z \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \\ 0 \end{pmatrix} \},$$

### then

- (A) the dimension of V equals 2
- (B) the dimension of V equals 4
- (C) the dimension of W equals 1
- (D)  $V \cap W = \{(0,0,0)\}$

# 18. Let M be a $3 \times 3$ non-zero, skew-symmetric real matrix. If I is the $3 \times 3$ identity matrix, then

- (A) M is invertible
- (B) The matrix I + M is invertible
- (C) There exists a non-zero real number  $\alpha$  such that  $\alpha I + M$  is not invertible
- (D) All the eigenvalues of M are real

## 19. Let X be a random variable with the moment generating function

$$M_X(t) = \frac{6}{\pi^2} \sum_{n>1} \frac{e^{t^2/n}}{n^2}, \quad t \in \mathbb{R}.$$

Then  $P(X \in \mathbb{Q})$ , where  $\mathbb{Q}$  is the set of rational numbers, equals

- (A) 0
- (B)  $\frac{1}{4}$
- (C)  $\frac{1}{2}$
- (D)  $\frac{3}{4}$

# 20. Let $\boldsymbol{X}$ be a discrete random variable with the moment generating function

$$M_X(t) = \frac{(1+3e^t)^2(3+e^t)^3}{1024}, \quad t \in \mathbb{R}.$$

Then

(A) 
$$E(X) = \frac{9}{4}$$

- (B)  $Var(X) = \frac{15}{32}$
- (C)  $P(X \ge 1) = \frac{27}{1024}$
- (D)  $P(X=5) = \frac{3}{1024}$

# 21. Let $\{X_n\}_{n\geq 1}$ be a sequence of independent random variables with $X_n$ having the probability density function as

$$f_n(x) = \begin{cases} \frac{1}{2n^{1/2}\Gamma(\frac{5}{2})} e^{-x^2/2}, & x > 0, \\ 0, & \text{otherwise.} \end{cases}$$

**Then** 

$$\lim_{n \to \infty} \left[ P(X_n > \frac{3n}{4}) + P(X_n > n + 2\sqrt{2n}) \right]$$

equals

- (A)  $1 + \Phi(2)$
- **(B)**  $1 \Phi(2)$
- (C)  $\Phi(2)$
- (D)  $2 \Phi(2)$

# 22. Let X be a Poisson random variable with mean $\frac{1}{2}$ . Then E((X+1)!) equals

- (A)  $2e^{-\frac{1}{2}}$
- **(B)**  $4e^{-\frac{1}{2}}$
- (C)  $4e^{-1}$
- (D)  $2e^{-1}$

# 23. Let X be a standard normal random variable. Then $P(X^3-2X^2-X+2>0)$ equals

- (A)  $2\Phi(1) 1$
- **(B)**  $1 \Phi(2)$
- (C)  $2\Phi(1) \Phi(2)$

(D) 
$$\Phi(2) - \Phi(1)$$

## **24.** Let *X* and *Y* have the joint probability density function

$$f(x,y) = \begin{cases} 2, & 0 \le x \le y \le 1, \\ 0, & \text{otherwise.} \end{cases}$$

Let  $a=E(Y|X=\frac{1}{2})$  and  $b=\text{Var}(Y|X=\frac{1}{2}).$  Then (a,b) is

- (A)  $(\frac{3}{4}, \frac{7}{12})$
- (B)  $(\frac{1}{4}, \frac{7}{12})$
- (C)  $\left(\frac{3}{4}, \frac{1}{48}\right)$
- (D)  $\left(\frac{3}{4}, \frac{1}{48}\right)$

# **25.** Let X and Y have the joint probability mass function

$$P(X = m, Y = n) = \frac{m+n}{21}, \quad m = 1, 2, 3; n = 1, 2,$$
 otherwise.

Then P(X=2|Y=2) equals

- (A)  $\frac{1}{3}$
- (B)  $\frac{2}{3}$
- (C)  $\frac{1}{2}$
- (D)  $\frac{1}{4}$

# 26. Let X and Y be two independent standard normal random variables. Then the probability density function of $Z=\frac{|X|}{|Y|}$ is

- (A)  $f(z) = \frac{1}{\sqrt{\pi}}e^{-z^2/2}, \quad z > 0, \quad 0,$  otherwise
- (B)  $f(z) = \frac{2}{\sqrt{2\pi}}e^{-z^2/2}, \quad z > 0, \quad 0, \quad \text{otherwise}$
- (C)  $f(z) = e^{-z}$ , z > 0, otherwise
- (D)  $f(z) = \frac{2}{\pi(1+z^2)}, \quad z > 0, \quad 0, \quad \text{otherwise}$

27. Let X and Y have the joint probability density function

$$f(x,y) = \begin{cases} e^{-y}, & 0 < x < y < \infty, \\ 0, & \text{otherwise.} \end{cases}$$

Then the correlation coefficient between X and Y equals

- (A)  $\frac{1}{3}$
- (B)  $\frac{1}{\sqrt{3}}$
- (C)  $\frac{1}{\sqrt{2}}$
- (D)  $\frac{2}{\sqrt{3}}$

28. Let  $x_1 = -2$ ,  $x_2 = 1$  and  $x_3 = -1$  be the observed values of a random sample of size three from a discrete distribution with the probability mass function

$$f(x;\theta) = P(X = x) = \begin{cases} \frac{1}{2\theta+1}, & x \in \{-\theta, -\theta+1, \dots, 0, \dots, \theta\}, \\ 0, & \text{otherwise,} \end{cases}$$

where  $\theta \in \{1, 2, \dots\}$  is the unknown parameter. Then the method of moment estimate of  $\theta$  is

- (A) 1
- (B) 2
- (C) 3
- (D) 4

**29.** Let X be a random sample from a discrete distribution with the probability mass function

$$f(x;\theta) = P(X = x) = \begin{cases} \frac{1}{\theta}, & x = 1, 2, \dots, \theta, \\ 0, & \text{otherwise,} \end{cases}$$

where  $\theta \in \{20, 40\}$  is the unknown parameter. Consider testing

$$H_0: \theta = 40$$
 against  $H_1: \theta = 20$ 

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at a level of significance  $\alpha=0.1$ . Then the uniformly most powerful test rejects  $H_0$  if and only if

- (A)  $X \leq 4$
- (B) X > 4
- (C)  $X \ge 3$
- (D) X < 3

# 30. Let $X_1$ and $X_2$ be a random sample of size 2 from a discrete distribution with the probability mass function

$$f(x;\theta) = P(X=x) = \begin{cases} \theta, & x = 0, \\ 1 - \theta, & x = 1, \end{cases}$$

where  $\theta \in \{0.2, 0.4\}$  is the unknown parameter. For testing

$$H_0: \theta = 0.2$$
 against  $H_1: \theta = 0.4$ ,

consider a test with the critical region

$$C = \{(x_1, x_2) \in \{0, 1\}^2 : x_1 + x_2 < 2\}.$$

Let  $\alpha$  and  $\beta$  denote the probability of Type I error and power of the test, respectively. Then  $(\alpha, \beta)$  is

- (A) (0.36, 0.74)
- **(B)** (0.64, 0.36)
- (C) (0.05, 0.64)
- (D) (0.05, 0.36)

# 31. Let $\{a_n\}_{n\geq 1}$ be a sequence of real numbers such that

$$a_n = \sum_{k=n+1}^{2n} \frac{1}{k}, \quad n \ge 1.$$

Then which of the following statement(s) is (are) true?

- (A)  $\{a_n\}_{n\geq 1}$  is an increasing sequence
- (B)  $\{a_n\}_{n\geq 1}$  is bounded below
- (C)  $\{a_n\}_{n\geq 1}$  is bounded above
- (D)  $\{a_n\}_{n\geq 1}$  is a convergent sequence

# 32. Let $\sum_{n\geq 1} a_n$ be a convergent series of positive real numbers. Then which of the following statement(s) is (are) true?

- (A)  $\sum_{n>1} (a_n)^2$  is always convergent
- (B)  $\sum_{n>1} \sqrt{a_n}$  is always convergent
- (C)  $\sum_{n\geq 1} \frac{\sqrt{a_n}}{n}$  is always convergent
- (D)  $\sum_{n\geq 1} \frac{\sqrt{a_n}}{n^{1/4}}$  is always convergent

# **33.** Let $\{a_n\}_{n\geq 1}$ be a sequence of real numbers such that $a_1=3$ and, for $n\geq 1$ ,

$$a_{n+1} = \frac{a_n^2 - 2a_n + 4}{2}.$$

### Then which of the following statement(s) is (are) true?

- (A)  $\{a_n\}_{n\geq 1}$  is a monotone sequence
- (B)  $\{a_n\}_{n\geq 1}$  is a bounded sequence
- (C)  $\{a_n\}_{n\geq 1}$  is convergent
- (D)  $a_n \to 2$  as  $n \to \infty$

### **34.** Let $f: \mathbb{R} \to \mathbb{R}$ be defined by

$$f(x) = \begin{cases} x^2(2 + \sin\frac{1}{x}), & x \neq 0, \\ 0, & x = 0. \end{cases}$$

#### Then which of the following statement(s) is (are) true?

(A) f attains its minimum at 0

- (B) f is monotone
- (C) f is differentiable at 0
- (D)  $f(x) > 2x^4 + x^3$ , for all x > 0

35. Let P be a probability function that assigns the same weight to each of the points of the sample space  $\Omega = \{1, 2, 3, 4\}$ . Consider the events

$$E = \{1, 2\}, \quad F = \{1, 3\}, \quad G = \{3, 4\}.$$

Then which of the following statement(s) is (are) true?

- (A) E and F are independent
- (B) E and G are independent
- (C) F and G are independent
- (D) E, F, and G are independent

36. Let  $X_1, X_2, \dots, X_n$ , where  $n \ge 5$ , be a random sample from a distribution with the probability density function

$$f(x;\theta) = \begin{cases} e^{-(x-\theta)}, & x \ge \theta, \\ 0, & \text{otherwise,} \end{cases}$$

where  $\theta \in \mathbb{R}$  is the unknown parameter. Then which of the following statement(s) is (are) true?

- (A) A 95% confidence interval of  $\theta$  has to be of finite length
- (B)  $\left(\min(X_1, X_2, \dots, X_n) + \frac{1}{n}\ln(0.05), \min(X_1, X_2, \dots, X_n)\right)$  is a 95% confidence interval of  $\theta$
- (C) A 95% confidence interval of  $\theta$  can be of length 1
- (D) A 95% confidence interval of  $\theta$  can be of length 2

**37.** Let  $X_1, X_2, \ldots, X_n$  be a random sample from  $U(0, \theta)$ , where  $\theta > 0$  is the unknown parameter. Let

$$X_{(n)} = \max(X_1, X_2, \dots, X_n).$$

Then which of the following is (are) consistent estimator(s) of  $\theta^3$ ?

- (A)  $8X_{(n)}^3$
- (C)  $\frac{2}{n} \sum_{i=1}^{n} X_i^3$ (D)  $\frac{nX_{(n)}^3 + 1}{n+1}$

38. Let  $X_1, X_2, \dots, X_n$  be a random sample from a distribution with the probability density function

$$f(x;\theta) = \begin{cases} c(\theta)e^{-x-\theta}, & x \ge \theta, \\ 0, & \text{otherwise,} \end{cases}$$

where  $\theta \in \mathbb{R}$  is the unknown parameter. Then which of the following statement(s) is (are) true?

- (A) The maximum likelihood estimator of  $\theta$  is  $\min(X_1, X_2, \dots, X_n)$
- (B)  $c(\theta) = 1$ , for all  $\theta \in \mathbb{R}$
- (C) The maximum likelihood estimator of  $\theta$  is  $\min(X_1, X_2, \dots, X_n)$
- (D) The maximum likelihood estimator of  $\theta$  does not exist

**39.** Let  $X_1, X_2, \dots, X_n$  be a random sample from a distribution with the probability density function

$$f(x;\theta) = \begin{cases} \theta^2 x e^{-\theta x}, & x > 0, \\ 0, & \text{otherwise,} \end{cases}$$

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where  $\theta>0$  is the unknown parameter. If  $Y=\sum_{i=1}^n X_i$ , then which of the following statement(s) is (are) true?

(A) Y is a complete sufficient statistic for  $\theta$ 

- (B)  $\frac{2n}{V}$  is the uniformly minimum variance unbiased estimator of  $\theta$
- (C)  $\frac{2n-1}{V}$  is the uniformly minimum variance unbiased estimator of  $\theta$
- (D)  $\frac{2n+1}{Y}$  is the uniformly minimum variance unbiased estimator of  $\theta$

# **40.** Let $X_1, X_2, \dots, X_n$ be a random sample from $U(\theta, \theta + 1)$ , where $\theta \in \mathbb{R}$ is the unknown parameter. Let

$$U = \max(X_1, X_2, \dots, X_n)$$
 and  $V = \min(X_1, X_2, \dots, X_n)$ .

### Then which of the following statement(s) is (are) true?

- (A) U is a consistent estimator of  $\theta$
- (B) V is a consistent estimator of  $\theta$
- (C) 2U V is a consistent estimator of  $\theta$
- (D) 2U V is a consistent estimator of  $\theta + 1$

# **41.** Let $\{a_n\}_{n\geq 1}$ be a sequence of real numbers such that

$$a_n = \frac{1+3+5+\dots+(2n-1)}{n!}, \quad n \ge 1.$$

**Then**  $\sum_{n\geq 1} a_n$  converges to ......

### 42. Let

$$S = \{(x, y) \in \mathbb{R}^2 : x \ge 0, \sqrt{4 - (x - 2)^2} \le y \le \sqrt{9 - (x - 3)^2} \}.$$

Then the area of S equals ......

### 43. Let

$$S = \{(x, y) \in \mathbb{R}^2 : |x| + |y| \le 1\}.$$

Then the area of S equals ......

## 44. Let

$$J = \frac{1}{\pi} \int_0^1 t^{-\frac{1}{2}} (1-t)^{\frac{3}{2}} dt.$$

Then the value of J equals .....

46. Let X and Y be two positive integer-valued random variables with the joint probability mass function

$$P(X = m, Y = n) = g(m)h(n), \quad m, n \ge 1,$$

where  $g(m) = \left(\frac{1}{2}\right)^{m-1}$ ,  $m \ge 1$ , and  $h(n) = \left(\frac{1}{3}\right)^n$ ,  $n \ge 1$ . Then E(XY) equals .......

47. Let E, F and G be three events such that

$$P(E \cap F \cap G) = 0.1$$
,  $P(G \mid F) = 0.3$  and  $P(E \mid F \cap G) = P(E \mid F)$ .

**Then**  $P(G \mid E \cap F)$  equals .....

**48.** Let  $A_1, A_2$  and  $A_3$  be three events such that

$$P(A_i) = \frac{1}{3}, \quad i = 1, 2, 3; \quad P(A_i \cap A_j) = \frac{1}{6}, \quad 1 \le i \ne j \le 3 \quad \text{and} \quad P(A_1 \cap A_2 \cap A_3) = \frac{1}{6}.$$

Then the probability that none of the events  $A_1, A_2, A_3$  occur equals ......

49. Let  $X_1, X_2, \dots, X_n$  be a random sample from the distribution with the probability density function

$$f(x) = \frac{1}{4}e^{-|x-4|} + \frac{1}{4}e^{-|x-6|}, \quad x \in \mathbb{R}.$$

**Then**  $\frac{1}{n} \sum_{i=1}^{n} X_i$  converges in probability to ......

50. Let  $x_1 = 1.1, x_2 = 2.2, x_3 = 3.3$  be the observed values of a random sample of size three from a distribution with the probability density function

$$f(x;\theta) = \begin{cases} \frac{1}{\theta}e^{-x/\theta}, & x > 0, \\ 0, & \text{otherwise,} \end{cases}$$

where  $\theta \in \{1, 2, \dots\}$  is the unknown parameter. Then the maximum likelihood estimate of  $\theta$  equals .......

**51.** Let  $f: \mathbb{R} \to \mathbb{R}$  be a differentiable function such that f' is continuous on  $\mathbb{R}$  with f'(3) = 18. Define

$$g_n(x) = n\left(f\left(x + \frac{5}{n}\right) - f\left(x - \frac{2}{n}\right)\right).$$

**Then**  $\lim_{n\to\infty} g_n(3)$  equals .....

**52.** Let  $M = \sum_{i=1}^{4} X_i X_i^T$ , where

$$X_1^T = [1 \ -1 \ 1 \ 0], \quad X_2^T = [1 \ 1 \ 0 \ 1], \quad X_3^T = [1 \ 3 \ 1 \ 0] \quad \text{and} \quad X_4^T = [1 \ 1 \ 1 \ 0].$$

Then the rank of M equals .....

## **53.** Let $f: \mathbb{R} \to \mathbb{R}$ be a differentiable function with

$$\lim_{x \to \infty} f(x) = \infty$$
 and  $\lim_{x \to \infty} f'(x) = 2$ .

**Then** 

$$\lim_{x \to \infty} \left( 1 + \frac{f(x)}{x^2} \right) \text{ equals .....}$$

54. The value of

$$\int_0^{\pi} \left( \int_0^x e^{\sin y} \sin x \, dy \right) dx$$

equals .....

## 55. Let X be a random variable with the probability density function

$$f(x) = \begin{cases} 4x^k, & 0 < x < 1, \\ x - \frac{x^2}{2}, & 1 \le x < 2, \\ 0, & \text{otherwise,} \end{cases}$$

where k is a positive integer. Then

$$P\left(\frac{1}{2} \le X \le \frac{3}{2}\right)$$
 equals .....

# 56. Let X and Y be two discrete random variables with the joint moment generating function

$$M_{X,Y}(t_1, t_2) = \left(\frac{1}{3}e^{t_1} + \frac{2}{3}\right)^2 \left(\frac{2}{3}e^{t_2} + \frac{1}{3}\right)^3, \quad t_1, t_2 \in \mathbb{R}.$$

**Then** P(2X + 3Y > 1) equals ......

# 57. Let $X_1, X_2, X_3$ and $X_4$ be i.i.d. discrete random variables with the probability mass function

$$P(X_1 = n) = \frac{3n-1}{4n}, \quad n = 1, 2, \dots,$$

**Then**  $P(X_1 + X_2 + X_3 + X_4 = 6)$  equals .....

### 58. Let X be a random variable with the probability mass function

$$P(X = n) = \frac{1}{10}, \quad n = 1, 2, \dots, 10.$$

**Then**  $E(\max\{X,5\})$  equals ......

# **59.** Let X be a sample observation from $U(\theta,\theta^2)$ distribution, where $\theta \in \{2,3\}$ is the unknown parameter. For testing

$$H_0: \theta = 2$$
 against  $H_1: \theta = 3$ ,

let  $\alpha$  and  $\beta$  be the size and power, respectively, of the test that rejects  $H_0$  if and only if  $X \geq 3.5$ . Then  $\alpha + \beta$  equals ......

## **60.** A fair die is rolled four times independently. For i = 1, 2, 3, 4, define

$$Y_i = \begin{cases} 1, & \text{if 6 appears in the } i\text{-th throw,} \\ 0, & \text{otherwise.} \end{cases}$$

**Then**  $P(\max\{Y_1, Y_2, Y_3, Y_4\} = 1)$  equals ......